

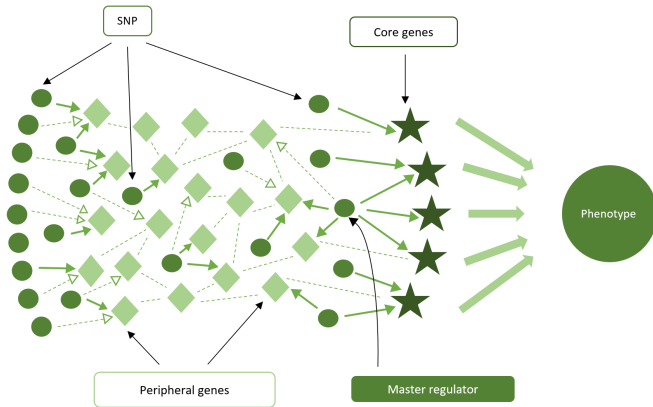
# Decoding the Genomic Network: A Dependence-Adjusted Approach for Functional Analysis of Omnigenic Models

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# Motivation : Omnigenic Network



---> Weak effect from trans SNP

---> Strong effect from cis SNP

--- Peripheral genes affecting each other/  
core genes

## Motivation : Why consider trans effects?

$$\begin{aligned}
 \text{Phenotypic variance} = & \underbrace{\sum_{j=1}^M \mu_j^2 V_{j;\text{cis}}}_{M \text{ core terms}} + \underbrace{\sum_{j=1}^M \mu_j^2 V_{j;\text{trans}}}_{M \text{ trans terms}} \\
 & + \underbrace{\sum_{j=1}^M \sum_{k=1}^j 2\mu_j \mu_k C_{jk}}_{M^2 - M \text{ covariance terms that mainly arise due to trans effect}} \quad \text{where } M^2 - M \geq M.
 \end{aligned}$$

# Statistical Modeling of the Biological Problem

- Suppose there are  $p$  SNPs in total across the entire genome.
- $Z_j$  is the summary-statistic for the  $j^{\text{th}}$  SNP.
- Suppose, WLG,  $S_0 = \{1, 2, \dots, p-2\}$  is the set of influential SNPs for a particular gene.
- $Z_j | \mu_j = \mu_j \mathbb{1}\{j \in S_0\} + \delta_j$ , where

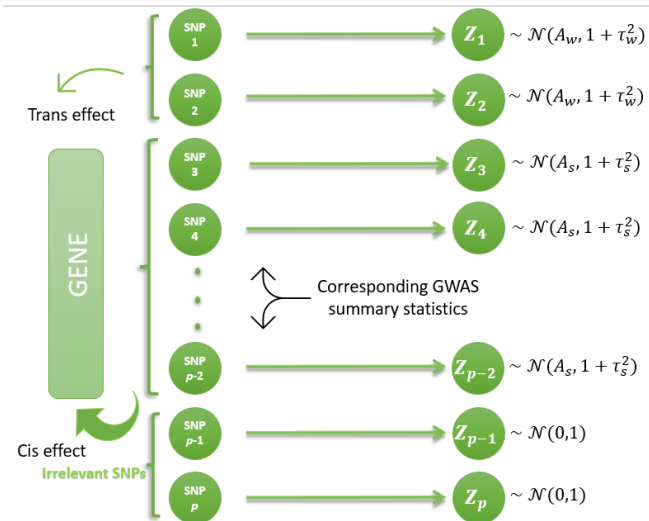
$$\delta_j \sim \mathcal{N}(0, 1) \text{ and } \mu_j \sim (1 - \epsilon)\delta_0 + \epsilon\mathcal{N}(A, \tau^2).$$

- Marginally,

$$Z_j \sim \begin{cases} (1 - \epsilon)\mathcal{N}(0, 1) + \epsilon\mathcal{N}(A, 1 + \tau^2), & \text{if } j \in S_0, \\ \mathcal{N}(0, 1), & \text{if } j \notin S_0. \end{cases}$$

- $A$  and  $\tau$  differ based on different kinds influential SNPs.

# One Gene System



# Model

- The formal statistical hypothesis then looks like

$$H_0 : Z_j \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1) \quad \text{vs.}$$

$$H_1 : Z_j \stackrel{\text{iid}}{\sim} (1 - \epsilon)\mathcal{N}(0, 1) + \epsilon\mathcal{N}(A, 1 + \tau^2), \quad \forall j = 1, \dots, p.$$

- HC and BJ tests can be used to determine how much each element of the vector of test-statistics digresses from  $\mathcal{N}(0, 1)$ .

HC

$$\max_{1 \leq j \leq p} \frac{\sqrt{p(j/p - p_{(j)})}}{\sqrt{p_{(j)}(1 - p_{(j)})}}$$

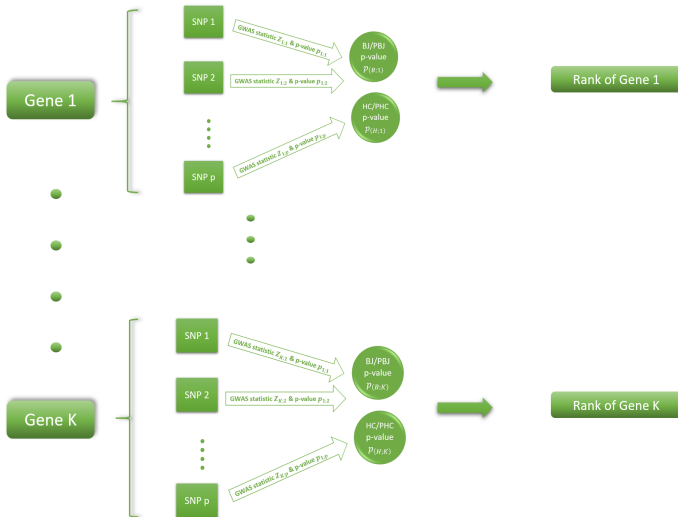
$$\sup_{t \in \mathbb{R}} \frac{\sqrt{p(\overline{F}_p(t) - \overline{\Phi}(t))}}{\sqrt{(\overline{\Phi}(t)(1 - \overline{\Phi}(t)))}}$$

BJ

$$\max_{1 \leq j \leq p} \frac{j}{p} \log \left( \frac{j/p}{p_{(j)}} \right) + (1 - \frac{j}{p}) \log \left( \frac{1 - j/p}{1 - p_{(j)}} \right)$$

$$\sup_{t \in \mathbb{R}} F_p(t) \log \left( \frac{F_p(t)}{\Phi(t)} \right) + \overline{F}_p(t) \log \left( \frac{\overline{F}_p(t)}{\overline{\Phi}(t)} \right)$$

# Ranking Scheme



## Dependence Adjustment

- For a given gene, let  $\Sigma$  be the covariance matrix of the  $p$  SNPs within the gene.
- Then, the spectral decomposition of  $\Sigma$  can be split into two parts as follows:

$$\Sigma = \sum_{m=1}^k \lambda_m \gamma_m \gamma_m^\top + \sum_{m=k+1}^p \lambda_m \gamma_m \gamma_m^\top = \sum_{m=1}^k \mathbf{b}_m \mathbf{b}_m^\top + \mathbf{A}$$

- **PFA:**  $Z_j = \mu_j + \mathbf{b}_j^\top \mathbf{W} + K_j$ ;  $j = 1, 2, \dots, p$  where  $\mathbf{W} \sim \mathcal{N}_k(\mathbf{0}, \mathbf{I}_k) \perp (K_1, \dots, K_p)^\top \sim \mathcal{N}_p(\mathbf{0}, \mathbf{A})$ .
- The dependence-adjusted p-values are then given by

$$p_i^* = 2\Phi\left(-a_j\left(Z_j - \mathbf{b}_j^\top \hat{\mathbf{W}}\right)\right)$$

where  $a_j$  is the suitable normalizing factor and  $\hat{\mathbf{W}}$  is a suitable estimator of  $\mathbf{W}$ .



## Detection Boundary

- **Reparameterization:**  $(A, \epsilon) \rightarrow (h, \gamma)$ , where  $h$  is the intensity parameter and  $\gamma$  is the sparsity parameter.
- Define  $m(\tau) = \min\left(\frac{1+\tau^2}{4}, \frac{1}{1+\tau^2}\right)$ . Then, the **detection boundary** is defined for  $1/2 \leq \gamma < 1$  as

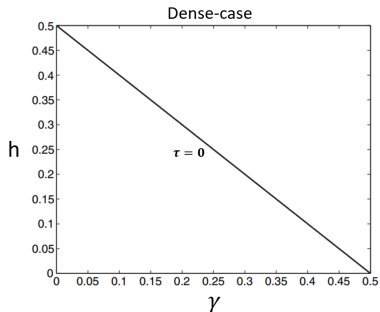
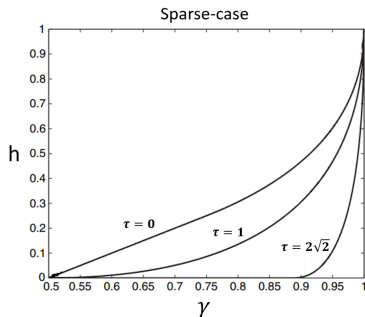
$$\rho(\gamma, \tau) = \begin{cases} (1 - (1 + \tau^2)\sqrt{1 - \gamma})^2 & , 1 - m(\tau) < \gamma < 1 \\ (1 - \tau^2)(\gamma - 1/2) & , 1/2 < \gamma < 1 - m(\tau) \end{cases}$$

and for  $0 \leq \gamma < 1/2$  as

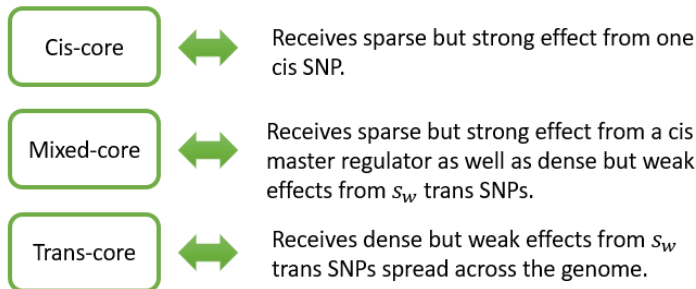
$$\rho(\gamma, \tau) = \begin{cases} 1/2 - \gamma & , \tau = 0 \\ \infty & , \tau > 0 \end{cases}$$

- The for the gene under consideration, HC and BJ tests attain **full asymptotic power** in the detectable region ( $h > \rho(\gamma, \tau)$ ) and the **type-I error vanishes** as  $n \rightarrow \infty$ .

# Detection Region

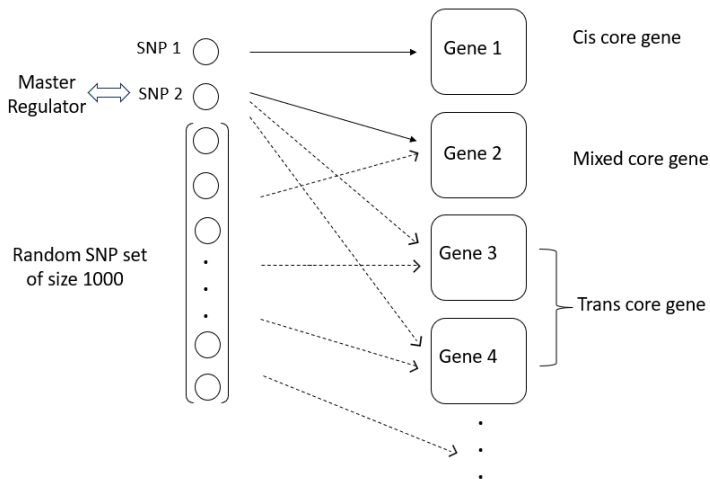


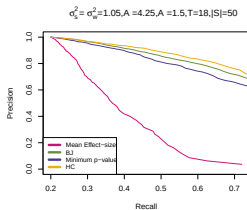
# Simulation Architecture



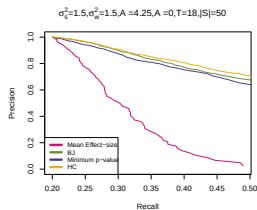
| Genes      | Index                 | Strong cis effect                                  | Weak trans effect                                        |
|------------|-----------------------|----------------------------------------------------|----------------------------------------------------------|
| Cis-core   | $k = 1$               | $\mu_{jk} \sim \mathcal{N}(A_s, \tau_s^2), j = 1,$ | -                                                        |
| Mixed-core | $k = 2$               | $\mu_{jk} \sim \mathcal{N}(A_s, \tau_s^2), j = 2$  | $\mu_{jk} \sim \mathcal{N}(A_w, \tau_w^2), j \in \tau_2$ |
| Trans-core | $k = 3, 4, \dots, 18$ | -                                                  | $\mu_{jk} \sim \mathcal{N}(A_w, \tau_w^2), j \in \tau_k$ |

# Simulation Setup

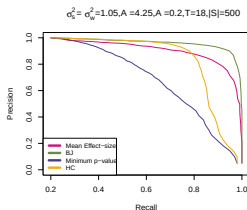




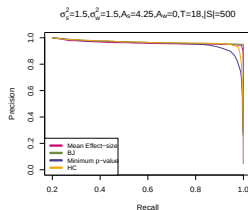
(a) Sparse Case 1



(b) Sparse Case 2



(c) Dense Case 1



(d) Dense Case 2

*Thank you!*